

## **Online Supplement**

Manuscript:

**“Anti-immigrant Prejudice in a Post-Socialist Context:  
The Role of Identity-based Explanations”**

## **S1. Anti-immigrant prejudice scale (Matić 2018; Matić, Löw, and Bratko 2019)**

### **a) Scale construction and validation**

In the process of construction and validation of the Anti-immigrant prejudice scale (Matić 2018), pilot study on the sample of 332 high-school students from the City of Zagreb and Zagreb County was conducted. The characteristics of the sample and the administration procedure were comparable to those in the main study. The initial pool of the items consisted of 24 statements denoting different expressions of prejudice toward immigrants. In the item selection process, several factors were considered: item content, descriptives, skewness and kurtosis, non-spurious item-total correlation, Cronbach's Alpha if item deleted. The items with higher variance and distribution closer to normal, as well as those which were positively and (relatively) highly correlated with the overall result on the scale and that did not (significantly) reduce the reliability of the scale were preferred. In addition to that, the statements that brought new information about the measured construct were preferred over those that were somewhat redundant to other items. The experience from the field trial was also taken into account, as to avoid statements that student found to be vague or ambiguous. Following the abovementioned criteria, twelve items were selected to form the final version of the scale (see section S1b below). Principal component analysis demonstrated clear unidimensionality of the final version of the Anti-immigrant prejudice scale, with only one eigenvalue exceeding 1. The extracted component accounted for 51.9 per cent of the data variance. All of the twelve items had high correlations with the latent dimension of anti-immigrant prejudice. Cronbach Alpha reliability of the scale was high ( $\alpha = .91$ ). Finally, bivariate correlations of anti-immigrant prejudice (measured by the final version of the scale) and several relevant constructs were inspected. As expected, the result on the Anti-immigrant prejudice scale correlated significantly and in the expected directions with other prejudice measures (with the prejudice toward gay men  $.59, p < .01$ ; with the prejudice toward individuals with mental illnesses  $.49, p < .01$ ; with the prejudice toward atheists  $.35, p < .01$ ; with the prejudice toward overweight people  $.34, p < .01$ ). The overall result on the newly constructed Anti-immigrant prejudice scale was also moderately associated with the two robust predictors of prejudice – right-wing authoritarianism and social dominance orientation ( $.38, p < .01$  and  $.43, p < .01$ , respectively).

### **b) Complete item list**

*Items marked with asterisk should be reverse coded.*

1. The arrival of a larger number of immigrants in Croatia should be prevented.
2. I sympathise with immigrants because of the problems they could experience in Croatia.\*
3. I would like to make friends with an immigrant.\*
4. Croats have a lot in common with immigrants.\*
5. If I looked for an employee in the future, I would give priority to a candidate from Croatia rather than to an immigrant.
6. If many immigrants settled in my neighbourhood, I would consider moving away.
7. If I had chance, I would help an immigrant to settle in Croatia.\*
8. I am afraid the presence of immigrants will lead to a weakened unity among Croatian people.

9. I do not like getting in contact with immigrants.
10. After moving to Croatia, immigrants should abandon their customs.
11. If I had chance, I would enjoy getting to know other cultures through contact with immigrants.\*
12. Our country can benefit from the cultural diversity of the population.\*

## S2. Descriptive statistics

Table S2. Descriptive statistics and Cronbach's Alpha values for the variables in the study ( $N = 1,034$ ).

Variable	<i>M</i>	<i>SD</i>	<i>Min.</i>	<i>Max.</i>	<i>α</i>
Political orientation	4.17	1.71	1	7	-
Cultural participation	1.87	0.71	1	4	.73
Reading habits	2.31	0.94	1	5	.57 <sup>a</sup>
Parental education	4.09	1.17	1	6	-
Conception of nationhood	2.87	0.65	1	4	.81
Perceived ethnic threat	3.07	0.66	1	5	.67
Anti-immigrant prejudice	3.01	0.87	1	5	.90

*Note.* <sup>a</sup> The low value of  $\alpha$  for the reading habits measure represents a limitation of the present study. However, it should be noted that the measure consists of only three items, in which case the Cronbach's alpha lower than 0.7 is acceptable (Hair et al., 2006).

### S3. Pearson correlations

Table S3. Pearson correlations between the variables in the study ( $N = 1,034$ ).

	1.	2.	3.	4.	5.	6.
1. Political orientation	-					
2. Cultural participation	-.13**	-				
3. Reading habits	-.12**	.37**	-			
4. Parental education	-.04	.28**	.10**	-		
5. Conception of nationhood	.20**	-.07*	-.06	-.10*	-	
6. Perceived ethnic threat	.23**	-.23**	-.13**	-.11**	.24**	-
7. Anti-immigrant prejudice	.27**	-.32**	-.22**	-.11**	.31**	.53**

Note. \*\*  $p < .01$ ; \*  $p < .05$ .

#### S4. Test of the measurement models

In the first step, we tested the measurement models based on the theoretical conceptualizations of the latent constructs using confirmatory factor analysis (CFA) in MPlus 7.0.

The one-factor model for the latent construct of anti-immigrant prejudice demonstrated a good fit ( $CFI = 0.98$ ;  $RMSEA = .04$ ;  $SRMR = .03$ ). The model is shown in *Figure S4a*.

The two-factor model for the latent construct of conception of nationhood demonstrated a poor fit ( $CFI = 0.89$ ;  $RMSEA = .10$ ;  $SRMR = .05$ ). The model is shown in *Figure S4b*. Therefore, we compared this model to the one-factor model using Akaike's Information Criterion (AIC) and Bayesian Information Criterion (BIC) – model fit indices for the non-nested model comparison. The smaller information criterion value indicated the better fit and the higher probability of replicating the model. The result of model fit comparison is shown in *Table S4*. The one-factor model showed a better fit (the smaller AIC and BIC values) and also a very good overall fit:  $CFI = 0.97$ ;  $RMSEA = .06$ ;  $SRMR = .03$ . The model is shown in *Figure S4c*.

The two-factor model for the latent construct of ethnic threat demonstrated a poor fit ( $CFI = 0.76$ ;  $RMSEA = .15$ ;  $SRMR = .07$ ). Therefore, we compared this model to the one-factor model using Akaike's Information Criterion (AIC) and Bayesian Information Criterion (BIC) – model fit indices for the non-nested model comparison. The smaller information criterion value indicated the better fit and the higher probability of replicating the model. The result of model fit comparison is shown in *Table S4*. The one-factor model showed a better fit (the smaller AIC and BIC values) and also an excellent overall fit:  $CFI = 1.00$ ;  $RMSEA = .02$ ;  $SRMR = .02$ . The model is shown in *Figure S4d*.

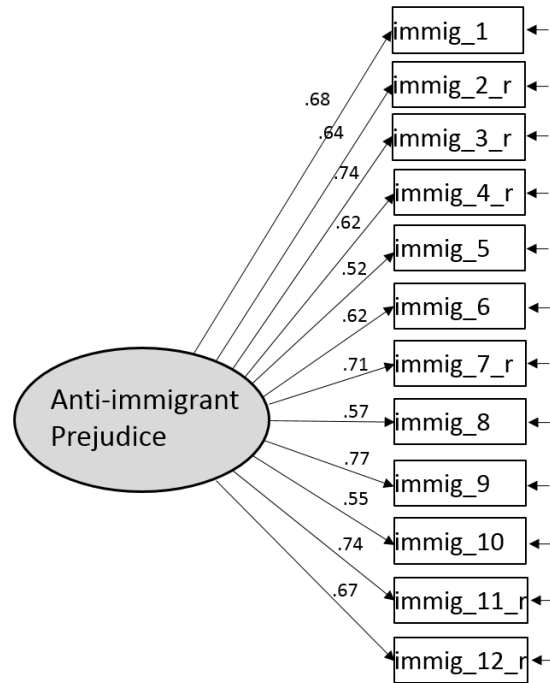
The two-factor model for the latent construct of cultural capital demonstrated a poor fit ( $CFI = 0.92$ ;  $RMSEA = .09$ ;  $SRMR = .05$ ). Therefore, we compared this model to the one-factor model using Akaike's Information Criterion (AIC) and Bayesian Information Criterion (BIC) – model fit indices for the non-nested model comparison. The smaller information criterion value indicated the better fit and the higher probability of replicating the model. The result of model fit comparison is shown in *Table S4*. The one-factor model showed a worse fit (the higher AIC and BIC values). Therefore, we specified a new model with two factors and a 2<sup>nd</sup> order factor and compared this model to the two-factor model also using the Akaike's Information Criterion (AIC) and Bayesian Information Criterion (BIC). The third model (*Table S4*) showed a better fit (the smaller AIC and BIC values) and also an excellent overall fit:  $CFI = 0.99$ ;  $RMSEA = .03$ ;  $SRMR = .02$ . Factor one of the cultural capital model included the cultural participation items and parental education, and was labelled as *cultural practices*. Factor two included the *reading habits* items; and was labelled accordingly. The second-order factor was labelled as *cultural capital*. The model is shown in *Figure S4e*.

Final (overall) measurement model fit the data well ( $\chi^2(465) = 1,067.02$ ;  $p < .001$ ;  $\chi^2 / df = 2.29$ ;  $CFI = 0.94$ ;  $RMSEA = .04$ ;  $SRMR = .05$ ).

Table S4. Model fit indices for the non-nested model comparisons.

Latent construct	Conception of nationhood		Ethnic threat		Cultural capital		
	1-factor	2-factor	1-factor	2-factor	1-factor	2-factor	2-factor with a 2 <sup>nd</sup> order factor
<i>AIC</i>	<b>19,527</b>	19,710	<b>15,013</b>	15,181	19,561	19,429	<b>19,361</b>
<i>BIC</i>	<b>19,663</b>	19,832	<b>15,124</b>	15,273	19,665	19,537	<b>19,479</b>

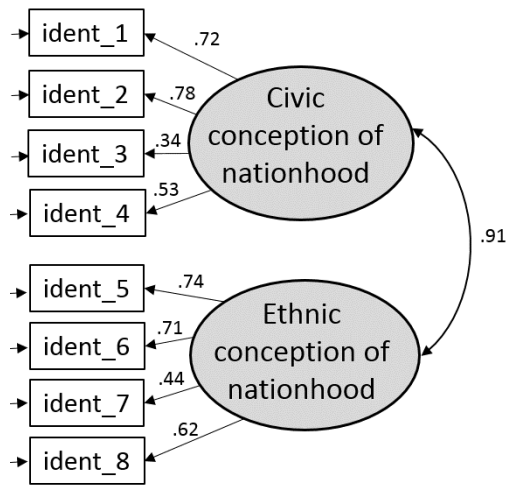
Note. *AIC* = Akaike information criterion, *BIC* = Bayesian information criterion. The smallest information criterion values for each latent construct are in bold.



*Figure S4a.* Final measurement model for the latent construct of anti-immigrant prejudice ( $N = 1,034$ ).

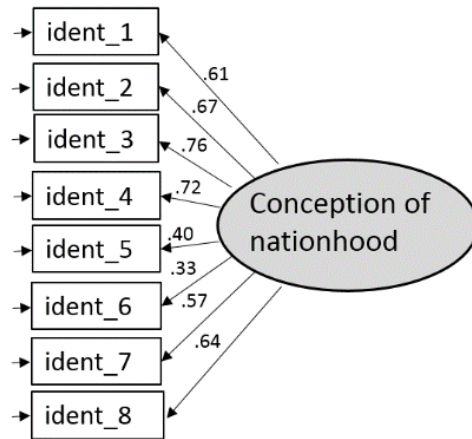
Note. Latent variables are shown in ovals, manifest variables in rectangles. Only standardized parameter estimates and significant paths are shown.





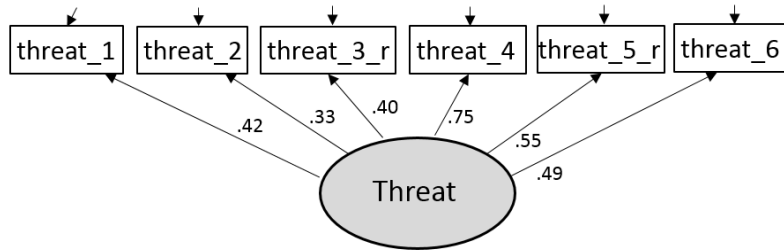
*Figure S4b.* Two-factor measurement model for the latent construct of conception of nationhood ( $N = 1,034$ ).

Note. Latent variables are shown in ovals, manifest variables in rectangles. Only standardized parameter estimates and significant paths are shown.



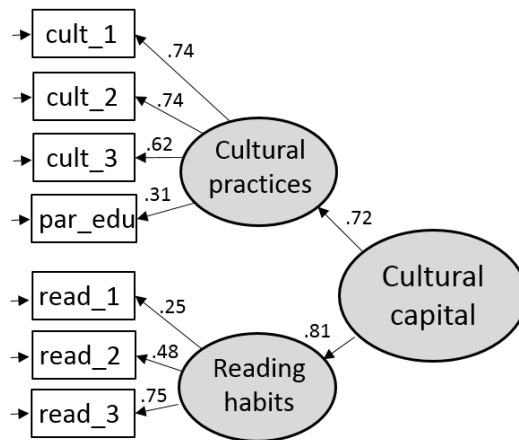
*Figure S4c.* Final measurement model for the latent construct of conception of nationhood ( $N = 1,034$ ).

Note. Latent variables are shown in ovals, manifest variables in rectangles. Only standardized parameter estimates and significant paths are shown.



*Figure S4d.* Final measurement model for the latent construct of ethnic threat ( $N = 1,034$ ).

Note. Latent variables are shown in ovals, manifest variables in rectangles. Only standardized parameter estimates and significant paths are shown.



*Figure S4e.* Final measurement model for the latent construct of cultural capital ( $N = 1,034$ ).

Note. Latent variables are shown in ovals, manifest variables in rectangles. Only standardized parameter estimates and significant paths are shown.

## S5. Structural model

In SEM analyses, cultural capital, conception of nationhood, ethnic threat and anti-immigrant prejudice were modeled as latent continuous variables with scale items as indicators, while single-item political orientation was modeled as an observed continuous variable.

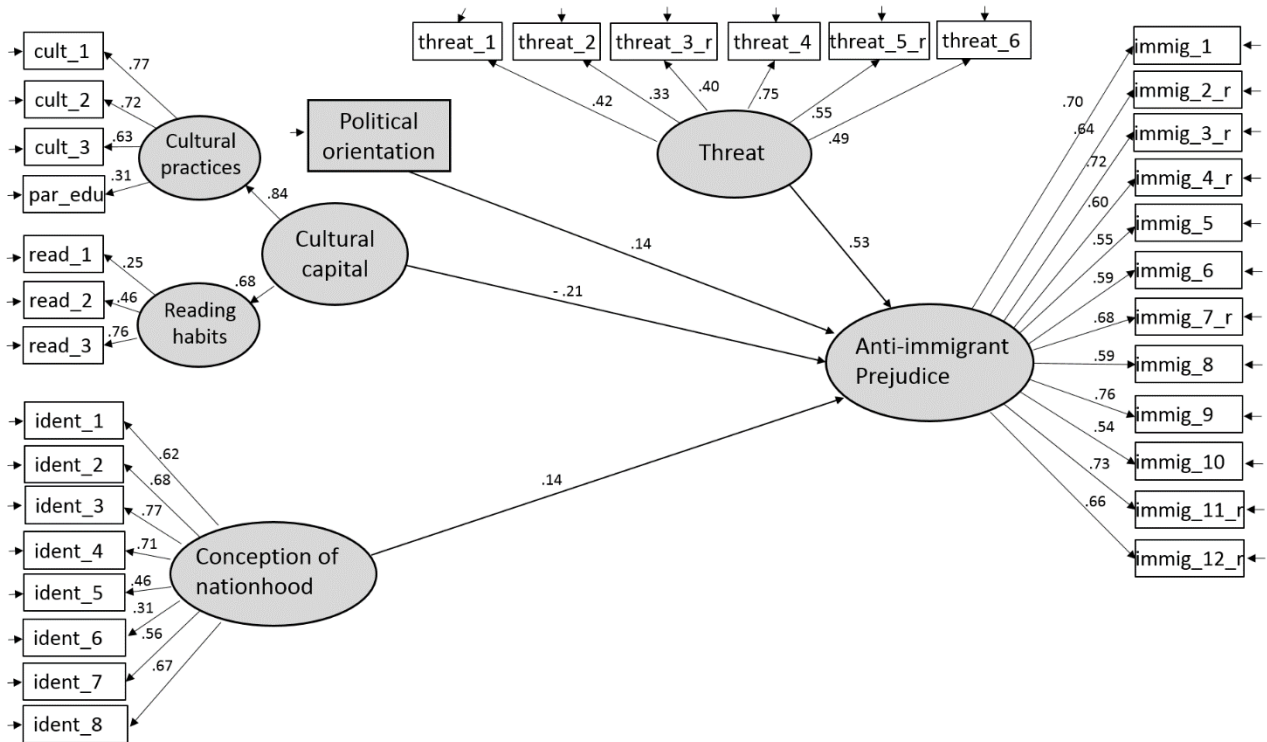


Figure S5. Structural model of relationships between identity-based variables and anti-immigrant prejudice ( $N = 1,034$ ).

Note. Latent variables are shown in ovals, manifest variables in rectangles. Only standardized parameter estimates and significant paths are shown.

*Table S5.* Unstandardized coefficients (*b*), standard errors (*s.e.*) and significance (*p*) for the structural model (Model A) and the final structural model with interaction of ethnic threat and political orientation (Model B).

Latent variable	Model A			Model B		
	<i>b</i>	<i>s.e.</i>	<i>p</i>	<i>b</i>	<i>s.e.</i>	<i>p</i>
Ethnic threat	1.146	0.264	.000	1.482	0.277	.000
Cultural capital	-0.580	0.213	.006	-0.563	0.207	.007
Political orientation	0.070	0.022	.001	0.075	0.022	.001
Conception of nationhood	0.186	0.062	.003	0.197	0.061	.001
Ethnic threat x Political orientation	-	-	-	-0.099	0.050	.047

## S6. Structural model with interactions

In this model, interactions were specified between: (a) latent continuous variables *ethnic threat* and *cultural capital*, (b) latent continuous variables *ethnic threat* and *conception of nationhood*, (c) latent continuous variable *ethnic threat* and observed continuous variable *political orientation*.

Table S6. Unstandardized coefficients (b) and 95% confidence intervals (95% CI) for interaction terms.

<b>Interaction term</b>	<b>b</b>	<b>95% CI for b</b>
Ethnic threat x Cultural capital	-0.089	[-0.721, 0.543]
Ethnic threat x Conception of nationhood	-0.109	[-0.320, 0.102]
Ethnic threat x Political orientation	-0.099*	[-0.180, -0.017]

Note. \*  $p < .05$